# **Original Article**

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# 24-Hour Urinary Sodium Excretion Association with Cardiovascular Events: A Systematic Review and Dose-Response Meta-Analysis

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# Abstract

**Objective** The relationship between sodium intake and cardiovascular (CV) events remains unconfirmed. Therefore, we carried out a systematic review and dose-response meta-analysis for evaluating the potential impact of 24-hour sodium excretion on CV risk. Besides, 24-hour sodium excretion was used to replace daily sodium diet intake.

**Methods** We searched ISI Web of Science, Embase, PubMed, and the Cochrane Library. Our study included cohort studies reporting hazard ratio (*HR*). The random-effects model was used for summarizing the total relative risks (*RRs*) between the included studies. In addition, the generalized least-squares regression was employed to fit the study model.

**Results** A total of 9 studies involving 645,006 participants were included in this study. A significant non-linear relationship was observed between sodium excretion and CV events ( $P_{non-linearity} < 0.001$ ). In studies collecting 24-h urine samples, the sodium excretion and CV events risk were associated linearly (*RR*: 1.04; 95% *CI*: 1.01, 1.07).

**Conclusion** In a linear dose-response manner, every 1 g increase in sodium intake was associated with an increased risk of CV events up to 4%. Further studies are required to validate our conclusions further.

**Key words:** Cardiovascular events; Meta-analysis; 24-h urinary sodium excretion; Dose-response; Sodium intake

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# INTRODUCTION

ccording to the World Health Statistics 2021, the disease burden has shifted to non-communicable diseases (NCDS) worldwide<sup>[1]</sup>. Among the four major NCDS. cardiovascular (CV) events' prevalence has contributed most to the associated mortality. Since 2000, CV events have increased by a quarter, reaching approximately 17.9 million in 2019<sup>[1]</sup>. Moreover, CV events have become a major public health problem due to their burden on health and

economic status on society. High blood pressure is one of the leading risk factors for CV events<sup>[2,3]</sup>. Several studies suggest that dietary sodium intake is closely related to blood pressure levels, and excessive sodium intake increases the risk for hypertension<sup>[4-7]</sup>. Therefore, the risk of CV events might be affected by sodium intake by altering the blood pressure levels.

The relationship between sodium intake and risk of CV events is not yet confirmed. Many studies were conducted to assess the association between sodium intake and the risk of CV events<sup>[8-14]</sup>.

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However, their conclusions were inconsistent. Some studies reported that the relationship between sodium intake and CV events was J-shaped or U-shaped, i.e., with the lower and higher sodium intake increases the risk of CV events<sup>[9,10,14]</sup>. Besides, some studies demonstrated that the sodium intakes were positively related to the risk of CV events<sup>[8,13]</sup>. In contrast, other studies suggested that the association between sodium intake and CV risk was not statistically significant<sup>[11,12]</sup>. Therefore, further investigation is required to understand the relationship between sodium intake and CV events.

Accurate assessment of dietary sodium intake is one of the major methodological challenges in dietary sodium research. Self-reported measurement methods of dietary intake, such as using food frequency questionnaire (FFQ) or single 24-h dietary recall, limited by errors in self-report and inaccuracies in food composition databases. The TRUE Consortium recommend against using self-reported methods for assessing dietary intake due to the inaccuracy of this method<sup>[15]</sup>. In healthy people, about 93 percent of sodium intake is excreted in the urine<sup>[16]</sup>. Twenty-four hour urinary excretion is considered as the gold standard method for assessing sodium intake<sup>[17]</sup>. To clarify the relationship between sodium intake and CV risk, we used 24-hour sodium excretion as a replacement for dietary intake, and conducted a systematic review and dose-response meta-analysis of published studies evaluating the association between 24-hour sodium excretion and CV risk. Furthermore, we discussed the limitations of current studies and provided hints for further studies.

#### METHODS

This meta-analysis was performed based on the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) statement (Supplementary Table S1 available in www.besjournal.com)<sup>[18]</sup>.

# Data Sources and Search Strategy

A systematic search was done on the scientific databases, including ISI Web of Science, Embase, PubMed, and the Cochrane Library, for extracting the relevant articles published from the inception to June 7, 2021. To explore the impact of 24-h urinary sodium excretion on CV events, we used a combination of keywords related to CV events and urinary sodium. Keywords associated with dietary sodium and potassium intake were also searched to avoid missing out on the studies that met our subject's criteria. In addition, the reference list was manually searched for the included studies. Detailed strategies for searching are shown in Supplementary Materials (available in www. besjournal.com).

## **Study Selection**

Two authors independently reviewed the studies, and any conflicts were resolved with the common consensus. First, the selection was conducted based on the title and abstract. Then the full texts of studies were reviewed for conforming our subject. Inclusion criteria: 1) full-texts in English; 2) cohort studies; 3) studies assessing the relationship of 24-h urinary sodium excretion with CV risk; 4) studies assessing 24-h urinary sodium excretion as exposure in at least three categories; 5) studies with the endpoints of either combined or single CV events (incidence and mortality); 6) studies providing adjusted effect estimates and 95% confidence intervals (*Cl*s).

Exclusion criteria: 1) non-English articles; 2) cross-sectional studies; 3) animal studies; 4) studies including kidney diseases patients; 5) studies with un-adjusted effect estimates. In addition, if multiple studies were from the same population, we included the longest follow-up study.

## Data Extraction

The following dataset was retrieved independently by the two authors: the first author's name, publication year, study region, study design, follow-up time, number of participants and cases, participant's gender, age at baseline, method of measuring 24-h urinary sodium excretion, assessment of exposure, outcome definition, covariates adjustment in the multivariate model, the indicator of effect estimates and other relevant characteristics. If the included studies reported models with different covariates adjustment, the meta-analysis selected the model with the most adjusted covariates.

## **Quality Assessment**

The Newcastle-Ottawa Quality Assessment Scale (NOS) was employed to estimate studies quality. The highest score was 9, and the lowest score was 0 for any study. Moreover, the studies that scored more than 7 were considered high quality.

### Statistical Analysis

The relative risks (*RR*s) with 95% *CIs* were used to calculate the effect size in studies, and then they were converted to natural logarithms to normalize

their distribution. The hazard ratios (*HR*s) were considered equal to the *RRs*. If the reference category was not the lowest, we took the lowest category as the reference and then recalculated the effect estimates<sup>[19]</sup>.

The association between 24-h urinary sodium excretion and the risk of CV events was evaluated via dividing sodium excretion into relatively low-level excretion and high-level excretion. Though the World Health Organization (WHO) recommended a sodium intake of < 2  $g/d^{[20]}$  and the Institute of Medicine (IOM) defined a tolerable upper sodium intake level of 2.3 g/d<sup>[21]</sup>, few people reached the recommended level. Global mean sodium intake was 3.95 g/d in 2010<sup>[22]</sup>, and the mean sodium excretion ranged from 3 to 4.9 g/d in the studies that included in this meta-analysis. Therefore, the cut-off point of 4 g was used in our meta-analysis, low-level excretion was defined as < 4 g and high-level excretion  $\geq$  4 g in our study. In one study, dose groups with medians higher than the cut-off point were classified as high-level, otherwise low-level. If studies reported sodium excretion in moles, it was converted into grams (1 mol = 23 g). In one study, the fixed-effects model combined the RRs of different dose groups. The random-effects model summarizes the total RRs between the included studies<sup>[23]</sup>. Sensitivity analysis was employed by removing one study to assess the potential effect of a single study on the pooled RR. Some studies included heart failure as an outcome, sensitivity analysis was conducted by excluding these studies. The study by Ma et al. was excluded due to duplication with our included population cohort, sensitivity analysis was also conducted by including this study<sup>[24]</sup>. The heterogeneity between the included studies was computed adopting Cochran's Q test and  $l^2$  statistic. Studies with values of  $l^2$ statistics less than 30% were considered to have no heterogeneity, whereas studies with more than 75% had notable heterogeneity or otherwise showed moderate heterogeneity<sup>[25]</sup>. To explain the observed heterogeneity, we performed subgroup analysis based on the follow-up duration, sample size, population information, method of urine collection, study quality, and adjusted confounders such as blood pressure or hypertension, potassium excretion. The funnel plot, Egger's test<sup>[26]</sup>, and Begg's test<sup>[27]</sup> were used to estimate the publication bias.

Then the dose-response meta-analysis was performed to explore the possible dose-response association, using the method proposed by Greenland, Longnecker, and Orsini<sup>[28,29]</sup>. The method

requires the number of cases and non-cases, or person-years in categories of urinary sodium excretion and the reported mean for every category of exposures. If the mean for each category of exposures was not reported, it was replaced with the median; otherwise, the midpoint in the exposure category was used. For the open category (highest and lowest), 20% high or low was calculated from the nearest cut-off point. Next, the non-linear relationship of urinary sodium excretion and CV events risk was estimated using the restricted cubic splines model with three knots at fixed percentiles (5%, 50%, and 95%)<sup>[30]</sup>. In contrast, the generalized least-squares regression was used to fit the model<sup>[29]</sup>. Finally, we pooled the effect estimates using the two-stage approach's random-effects models. Whenever the second spline coefficient was equal to zero, the non-linear P-value for the doseresponse meta-analysis was computed. The Stata 14.0 and Stata 16.0 software (StataCorp, College Station, Texas 77845 USA) were used for the statistical analysis.

## RESULTS

#### Literature Selection

A total of 9,495 records were searched by the online databases and the references of the relevant studies. After excluding the duplicates and examining studies by titles and abstracts, 90 articles were included for assessing the full text. According to the inclusion and exclusion criteria, finally, 9 studies were eventually included. Selection details are described in Figure 1.

#### Study Characteristics

The characteristics of the 9 included studies are shown in Table  $1^{\left[9,10,12,14,31\cdot35\right]}.$  Further information (the categories of urinary sodium excretion in original studies, the original effect size and transformed effect size, and covariates/factors adjusted in multivariate model) is shown in Supplementary Table S2 available in www.besjournal.com. The studies included in the meta-analysis were published between 2011 and 2021. A total of 645,006 participants with a median follow-up period of 4.7 to 19.1 years were recruited in our study. Two studies were carried out in the United States<sup>[10,12]</sup>, two in the United Kingdom<sup>[32,35]</sup>, one in the Netherlands<sup>[31]</sup>, one in Finland<sup>[33]</sup>, one in China<sup>[34]</sup>, and two multi-centric studies with recruited subjects from several countries<sup>[9,14]</sup>. For urinary sodium excretion

assessment, 4 studies collected the 24-h urine samples<sup>[10,12,31,33]</sup>, 2 collected the morning fasting urine samples<sup>[9,14]</sup>, 2 collected the spot urine samples<sup>[32,35]</sup>, whereas the 1 collected the overnight urine<sup>[34]</sup>. The mean score of included studies for NOS was 8 (Supplementary Table S3 available in www.besjournal.com). Moreover, most studies results were greater than  $7^{[10,31-35]}$ . One in nine studies reported data on men and women, respectively<sup>[35]</sup>. The covariates adjusted in most multivariate models were the baseline age, body mass index (BMI), ethnicity/race, sex, alcohol intake, smoking, blood pressure or hypertension, and blood fat level. The outcome definition of studies included total CVD, CVD death, heart failure, coronary heart disease, all caused mortality, atrial fibrillation, and atrial flutter. The daily urinary sodium excretion ranged from 1.265 to 9.4 g. Besides, all studies reported HR. Three studies reported a J-shaped relationship between calculated sodium excretion and risk of CV events<sup>[9,14,33]</sup>, 2 studies reported a Ushaped<sup>[32,35]</sup>, and 2 studies reported a linear association<sup>[10,34]</sup>. On the contrary, 2 studies suggested no significant effect of sodium excretion upon CV events<sup>[12,31]</sup>.

# Associations between Different Levels of Sodium Excretion and the Risk of CV Events

The association between the high and the low

levels of sodium excretion with the risk of CV events are shown in Figure 2. The total RR was 0.91 (95% CI: 0.86, 0.96;  $l^2 = 16.82\%$ ) and 1.01 (95% Cl: 0.91, 1.12;  $I^2$  = 84.25%) for the association between low-level sodium excretion (vs. the lowest of 1.265 g) and the high-level sodium excretion (vs. the lowest of 1.265 g) with the risk of CV events, respectively. However, significant heterogeneity between the studies was found (P = 0.01) in the high-level group. Therefore, subgroup analyses were carried out based on the follow-up duration, sample size, population information, method of urine collection, study quality, and adjusted confounders such as blood pressure or hypertension, potassium excretion (Table 2). Further, the subgroup of sample size and method of urine collection could explain the part of the heterogeneity.

# Dose-response Association between the Sodium Exposure and the Risk of CV Events

Based on the included studies, the doseresponse meta-analysis was carried out. A significant non-linear relationship was found between sodium excretion and the risk of CV events ( $P_{non-linearity} <$ 0.001) (Figure 3). In addition, the Wald test's result showed that the two slopes in the non-linear model were significantly different (P < 0.001). We also excluded 5 studies using spot urine samples and performed a dose-response meta-analysis with the



Figure 1. Flow chart of the literature search for studies.

Author, year	Country	Follow-up (years)	Subjects/ events ( <i>n</i> )	Age (y)	Gender	Methods of Na measurement	Study population	Outcome definition
Martin J. O' Donnell, 2011 <sup>[9]</sup>	40 countries	4.7	28,880/ 4,729	66.52 ± 7.22	men/ women	collected morning fasting urine samples	participants in ONTARGET and the TRANSCEND trials	composite outcome (CV mortality, myocardial infarction, stroke, and hospitalization for CHF)
Roman Pfister, 2013 <sup>[32]</sup>	the United Kingdom	12.9	19,857/ 1,210	58.0 ± 9.2	men/ women	collected casual urine specimen	general population	heart failure
Michel M. Joosten, 2013 <sup>[31]</sup>	the Netherlands	10.5 (IQR: 9.9-10.8)	7,543/ 452	28-75	men/ women	collected 24-h urine samples	general population	CHD was defined as incident cardiac morbidity and mortality (including myocardial infarction, acute and subacute ischemic heart disease, and coronary artery bypass grafting or percutaneous transluminal coronary angioplasty)
Nancy R. Cook, 2014 <sup>[10]</sup>	the United States	5	2,312/ 193	31–50	men/ women	collected 24-h urine samples	general population	CVD or CVD death, including myocardial infarction, stroke, coronary artery bypass graft, percutaneous transluminal coronary angioplasty, or death from cardiovascular disease
Pamela Singer, 2014 <sup>[12]</sup>	the United States	18.6	3,505/ 399	52.4 ± 9.9	men/ women	collected 24-h urine samples	participants were individuals in a union- sponsored, worksite hypertension program in New York City between 1978 and 1999	coronary artery disease, including MI, ischemic heart disease, heart failure, and hypertensive heart disease
Martin O' Donnell, 2019 <sup>[14]</sup>	628 urban and rural communities in low, middle, and high income countries	8.2	103,200/ 7,884	35-70	men/ women	collected morning fasting midstream urine sample	general population	the primary composite outcome was all caused mortality or myocardial infarction or stroke or heart failure
Matti A. Vuori, 2020 <sup>[33]</sup>	Finland	14	4,517/ 424	45.4 ± 11.4	men/ women	collected 24-h urine samples	general population	CVD was defined as the onset of CHD, stroke, or heart failure
Wuopio J, 2020 <sup>[35]</sup>	the United Kingdom	8.2 ± 1.0	215,535/ 3,751 257,545/ 2,221	40-69	men women	collected midstream spot urine samples	general population	atrial fibrillation or atrial flutter
Yi-Jie Wang, 2021 <sup>[34]</sup>	China	19.1 (IQR: 7.4-21.4)	2,112/ 279	54 ± 12	men/ women	recorded the sleep time and calculated the 24-h urine amount from sleep time and morning voiding urine	general population	Incident CVD events, including CHD and stroke

Table 1. Characteristics of the included studies in this Meta-Analysis

**Note.** 24 h UNaE: 24 h urinary sodium excretion; CV: cardiovascular; CHD: coronary heart disease; CVD: cardiovascular disease; CHF: congestive heart failure; ONTARGET: Ongoing Telmisartan Alone and in combination with Ramipril Global Endpoint Trial; TRANSCEND: Telmisartan Randomized Assessment Study in ACE Intolerant Subjects with Cardiovascular Disease.

remaining 4 studies. The results showed that the sodium excretion and risk of CV events were associated linearly (P = 0.02) (Figure 4). The *RR* in the model showed that every 1 g increase in sodium excretion was associated with an increase in the risk of CV events up to 4% (*RR*: 1.04; 95% *CI*: 1.01, 1.07).

None of the studies significantly affected the combined RR for sensitivity analysis in the low-level and high-level groups (Supplementary Figure S1 available in www.besjournal.com). In studies without an outcome of heart failure<sup>[10,31,34,35]</sup>, the combined *RR* for sensitivity analysis was 0.94 (95% *CI*: 0.86, 1.02) in the low-level group and 1.17 (95% *CI*: 0.91,

1.50) in the high-level group. When included in the study by Ma et al.<sup>[24]</sup>, the combined *RR* was 0.95 (95% *Cl*: 0.88, 1.04) in the low-level group and 1.07 (95% *Cl*: 0.94, 1.23) in the high-level group.

# **Publication Bias**

The results of the funnel plots did not exclude any possible publication bias for the low-level and high-level groups (Supplementary Figure S2 available in www.besjournal.com). Neither the Egger's nor Begg's test detected any evidence of publication bias for the low-level group (*P*-value for Egger: 0.132; *P*value for Begg: 0.252). The Egger's test showed

A	Study		<i>RR</i> with 95% <i>Cl</i>	Weight (%)
	O'Donnell, 2011 <sup>[9]</sup>		0.91 [0.81, 1.03]	14.37
	Pfister, 2013 <sup>[32]</sup>		0.78 [0.69, 0.89]	13.89
	Joosten, 2013 <sup>[31]</sup>		1.04 [0.86, 1.25]	7.02
	Cook, 2014 <sup>[10]</sup>		1.10 [0.54, 2.24]	0.57
	Singer, 2014 <sup>[12]</sup>		1.01 [0.79, 1.29]	4.48
	O'Donnell, 2019 <sup>[14]</sup>		0.91 [0.82, 1.00]	19.91
	Vuori, 2020 <sup>[33]</sup>		1.11 [0.79, 1.56]	2.41
	Wuopio, 2020 <sup>[35]</sup>	-	0.91 [0.85, 0.97]	33.15
	Wang, 2021 <sup>[34]</sup>		0.98 [0.76, 1.25]	4.22
	Overall	•	0.91 [0.86, 0.96]	
	Heterogeneity: $T^2 = 0.00$ , $I^2 = 16.82\%$ , $H^2 = 1.20$ Test of $\theta_i = \theta_j$ : $Q(8) = 9.91$ , $P = 0.27$ Test of $\theta = 0$ : $z = -3.43$ , $P = 0.00$			
	Random-effects REML model	1 2	-	
в	Study		<i>RR</i> with 95% <i>Cl</i>	Weight (%)
В	Study O'Donnell, 2011 <sup>[9]</sup>	•	<i>RR</i> with 95% <i>CI</i> 0.94 [0.85, 1.03]	Weight (%) 16.07
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup>	*	<i>RR</i> with 95% <i>CI</i> 0.94 [0.85, 1.03] 0.85 [0.75, 0.96]	Weight (%) 16.07 14.76
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup>	*	<i>RR</i> with 95% <i>CI</i> 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61]	Weight (%) 16.07 14.76 7.29
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup>	*	<i>RR</i> with 95% <i>CI</i> 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49]	Weight (%) 16.07 14.76 7.29 3.57
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup>	*	<i>RR</i> with 95% <i>CI</i> 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] –1.50 [0.91, 2.49] 1.00 [0.71, 1.41]	Weight (%) 16.07 14.76 7.29 3.57 6.20
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup>	*	RR with 95% CI 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup> Vuori, 2020 <sup>[33]</sup>	*	RR with 95% CI 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97] 1.22 [0.98, 1.52]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63 10.25
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup> Vuori, 2020 <sup>[33]</sup> Wuopio, 2020 <sup>[35]</sup>		RR with 95% CI 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97] 1.22 [0.98, 1.52] 0.92 [0.88, 0.97]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63 10.25 17.73
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup> Vuori, 2020 <sup>[33]</sup> Wuopio, 2020 <sup>[35]</sup> Wang, 2021 <sup>[34]</sup>		RR with 95% C/ 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97] 1.22 [0.98, 1.52] 0.92 [0.88, 0.97] 1.43 [1.02, 2.00]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63 10.25 17.73 6.50
В	Study   O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup> Vuori, 2020 <sup>[33]</sup> Wuopio, 2020 <sup>[35]</sup> Wang, 2021 <sup>[34]</sup> Overall		RR with 95% CI 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97] 1.22 [0.98, 1.52] 0.92 [0.88, 0.97] 1.43 [1.02, 2.00] 1.01 [0.91, 1.12]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63 10.25 17.73 6.50
В	Study O'Donnell, 2011 <sup>[9]</sup> Pfister, 2013 <sup>[32]</sup> Joosten, 2013 <sup>[31]</sup> Cook, 2014 <sup>[10]</sup> Singer, 2014 <sup>[12]</sup> O'Donnell, 2019 <sup>[14]</sup> Vuori, 2020 <sup>[33]</sup> Wuopio, 2020 <sup>[35]</sup> Wang, 2021 <sup>[34]</sup> Overall Heterogeneity: $T^2 = 0.02$ , $l^2 = 84.25\%$ , $H^2 = 6.35$ Test of $\theta_i = \theta_j$ : $Q(8) = 20.83$ , $P = 0.01$ Test of $\theta = 0$ : $z = 0.14$ , $P = 0.88$		RR with 95% Cl 0.94 [0.85, 1.03] 0.85 [0.75, 0.96] 1.19 [0.88, 1.61] -1.50 [0.91, 2.49] 1.00 [0.71, 1.41] 0.92 [0.88, 0.97] 1.22 [0.98, 1.52] 0.92 [0.88, 0.97] 1.43 [1.02, 2.00] 1.01 [0.91, 1.12]	Weight (%) 16.07 14.76 7.29 3.57 6.20 17.63 10.25 17.73 6.50

**Figure 2.** (A) Forest Plot of relative risks (*RRs*) and 95% Confidence Intervals (*CIs*) for the association between low-level sodium excretion (< 4 g/d) and the risk of CV events. (B) *RRs* and 95% *CIs* for the association between high-level sodium excretion ( $\geq$  4 g/d) and the risk of CV events.

-	No. of		Lov	v level			Higl	n level	
Subgroups	studies	Pooled <i>RR</i> (95% <i>Cl</i> )	P value	l² (%)	P value for heterogeneity	Pooled <i>RR</i> (95% <i>Cl</i> )	P value	l² (%)	<i>P</i> value for heterogeneity
Follow-up duration									
< 10 years	4	0.910 (0.866, 0.958)	0.0003	< 0.01	0.9611	0.926 (0.895, 0.957)	< 0.0001	0.01	0.2904
> 10 years	5	0.947 (0.821, 1.093)	0.4575	54.77	0.0481	1.093 (0.898, 1.329)	0.3759	68.49	0.0037
Sample size									
< 10,000	5	1.027 (0.912, 1.156)	0.6632	< 0.01	0.9789	1.228 (1.070, 1.410)	0.0035	< 0.01	0.5886
> 10,000	4	0.886 (0.836, 0.940)	0.0001	27.48	0.1930	0.918 (0.889, 0.948)	< 0.0001	0.02	0.6090
Conducted in gene	ral populati	ion							
Yes	7	0.912 (0.843, 0.986)	0.0207	42.63	0.1654	1.049 (0.902, 1.220)	0.5358	91.08	0.0021
No	2	0.932 (0.834, 1.042)	0.2162	< 0.01	0.4745	0.943 (0.861, 1.033)	0.2062	<0.01	0.7316
Method of urine co	ollection								
Collected 24-h urine samples	4	1.042 (0.911, 1.192)	0.5489	< 0.01	0.9721	1.190 (1.023, 1.385)	0.0242	< 0.01	0.6025
Collected spot urine samples	5	0.892 (0.847, 0.939)	< 0.0001	12.98	0.2670	0.922 (0.893, 0.952)	< 0.0001	< 0.01	0.0744
Study quality									
≤7	3	0.918 (0.852, 0.989)	0.0241	0.01	0.7224	0.928 (0.887, 0.971)	0.0012	< 0.01	0.8695
> 7	6	0.925 (0.826, 1.035)	0.1743	55.12	0.1034	1.094 (0.909, 1.316)	0.3439	84.39	0.0010
Adjusted blood pre	ssure or hy	pertension							
Yes	4	0.883 (0.815, 0.958)	0.0026	42.65	0.1638	0.921 (0.885, 0.959)	0.0001	< 0.01	0.0365
No	5	0.967 (0.876, 1.068)	0.5129	14.68	0.5758	1.084 (0.917, 1.281)	0.3444	59.57	0.0211
Adjusted potassiun	n excretion								
Yes	4	0.960 (0.873, 1.056)	0.4035	< 0.01	0.6694	1.047 (0.879, 1.247)	0.6076	40.92	0.1626
No	5	0.892 (0.829, 0.960)	0.0023	36.53	0.1534	0.999 (0.855, 1.168)	0.9923	92.71	0.0057

Table Z. Subgroup analysis of different levels of sodium excretion and the risk of CV evel	Table 2	2. Subgroup	analysis of	different levels	of sodium	excretion a	and the risk	of CV eve	nts
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**Figure 3.** Non-linear dose-response association between 24-h Sodium excretion and the risk of CV events. Dashed lines indicated 95% confidence intervals. Reference standard was 1.265 g/24 h.



**Figure 4.** Linear dose-response association between 24-h Sodium excretion and the risk of CV events. Dashed lines indicated 95% confidence intervals.

#### DISCUSSION

This current study included 645,006 participants; 21,542 had CV events. In the low-level group, the relationship between sodium excretion and the risk of CV events was statistically significant (P < 0.001). However, the relationship between sodium excretion and the risk of CV events was not statistically significant in the high-level group (P = 0.885). A significant non-linear relationship was observed between sodium excretion and CV risk ( $P_{non-linearity} < 0.001$ ). Moreover, the result of the dose-response meta-analysis with the 4 studies collecting 24-h urine samples showed that the sodium excretion and CV events risk were associated linearly. It is indicated that increased sodium excretion might be a risk factor for the CV events.

This meta-analysis showed significant heterogeneity in the high-level sodium excretion ( $l^2 = 84.25\%$ ). In the subgroup analysis of the method of urine collection, the pooled *RR* of studies with 24-h urine collection was 1.190 (1.023, 1.385), suggesting the statistically significant relationship of urinary sodium excretion with the risk of CV events. However, the result of studies using collected spot urine samples was converse. This might be due to the error in 24hour sodium excretion computed by collecting spot urine samples. Then the sodium excretion calculated by spot urine samples was not accurate and thus might have affected the authenticity of the final results.

Our study observed a J-shaped relationship between sodium intake (measured by urinary sodium biomarker) and CV events. The non-linear dose-response connection was also found in another meta-analysis<sup>[36]</sup>. In addition, the proof of a Ushaped association shown in a previous metaanalysis indicated that both low sodium and high sodium intakes were connected to the higher risk of CV events<sup>[37]</sup>. The RR in our linear model showed that every 1 g increase in sodium intake was associated with an increase in the risk of CV events up to 4%. Likewise, a recent meta-analysis showed that the sodium intake and risk of CV events were associated linearly. And that meta-analysis showed that every 1 g increase in sodium intake was associated with an increase in CV events risk up to 6%<sup>[38]</sup>. The reasons for the different results may be: 1) the methods of sodium measurement were different (urinary sodium biomarker or self-reported dietary intake measurement); 2) different definitions of study outcomes; 3) the ranges of sodium dose and cut-off points were different in the included studies.

Sodium is an essential nutrient component in our diet for maintaining the proper blood volume and blood pressure. The increase in sodium concentration influences the role of the reninangiotensin-aldosterone system and elevates the heart burden<sup>[39,40]</sup>. In addition, blunt renal salt excretion enlarges the extracellular fluid volume and increases blood pressure, manifested as the salt sensitivity of blood pressure<sup>[41]</sup>. Moreover, a high-sodium diet is related to myocardial contractility and changes in the proteins associated with calcium homeostasis<sup>[42]</sup>. Therefore, dietary sodium intake might influence the risk of CV events.

According to the linear dose-response metaanalysis conducted with studies collecting 24-h urine samples, increased sodium excretion was a risk factor for the CV events. We recommend that sodium intake be adapted based on individual risk factors. The public should be educated about the dangers of excessive sodium intake. Moreover, no convincing explanation exists for the mechanism of lower sodium intake related to a higher risk of CV events. Therefore, a non-pharmacological measure of reducing sodium intake is a cost-effective option for preventing CV events. There are also some studies examined the association between 24-h urinary potassium excretion and CV events<sup>[9,14,43,44]</sup>. The results of these studies were similar, that is, higher potassium excretion was associated with a lower risk of CV events. It is indicated that salt substitute with lower sodium and higher potassium may reduce the risk of CV events effectively.

In epidemiological studies without a urinary sodium biomarker, self-reported dietary intake measurement methods were considered significant bias<sup>[45]</sup>. The formulas for estimating the sodium intake from spot urine contain certain factors (e.g., sex, age, and creatinine concentration) strongly associated with CV events<sup>[46,47]</sup>. Although 24-hour urinary excretion is considered the gold standard method for assessing sodium intake<sup>[17]</sup>, a single measurement is not sufficient for providing a reliable estimation of long-term average sodium intake. We could not modify a single measurement's mean estimate of regression dilution bias<sup>[48]</sup>. Individuals at high risk of CV events might consciously consume less sodium, making the association between sodium intake and disease unreliable. Furthermore, nonadjustment of total energy intake can also produce errors because the total energy intake is highly correlated with sodium intake<sup>[49]</sup>. Moreover, almost all previous articles are observational studies, whereas randomized controlled trials (RCTs) are lacking, which are required to confirm the existing conclusions. To standardize the data reporting, we recommend (i) using multiple inconsecutive 24-hour urine collections; (ii) selecting subjects without high risk of CV events; (iii) adjusting sodium intake for the total energy intake.

This study had the following strengths. First, different meta-analysis approaches were employed to clarify the relationship between the 24-hour sodium excretion and CV events risk. Second, the total sample size of the meta-analysis was large (n =645,006), and multiple confounding factors that might affect the results were adjusted in the included studies to make the results more reliable. Third, we plotted a dose-response relationship graph, with the lowest point of the curve providing a threshold dose for sodium excretion that increases the risk of CV events. Finally, in all studies, the doseresponse analysis combined with different dose groups of sodium excretion included a wider range of sodium excretion than a single study, making the meta-analysis results more comprehensive and accurate.

However, the study had some limitations. First, multiple confounding factors adjusted in different included studies were not identical, which might have affected the authenticity of the results. Second, this meta-analysis assumed that the heterogeneity among the studies was significant due to the differences in study quality, sample size, follow-up duration, and division of sodium excretion dose groups. Third, several studies did not use the lowest dose group as the reference to calculate the effect estimates. Although we recalculated the effect estimates, it inevitably caused some errors in the recalculated 95% Cls. Finally, only four studies collected 24-hour urine samples to measure sodium excretion. The remaining five used samples of spot urine, which might have led to the inaccurate calculations of 24-hour sodium excretion.

### CONCLUSION

To conclude, this meta-analysis showed a significant relationship between the 24-h sodium excretion and the risk of CV events. In studies collecting 24-h urine samples, a linear relationship was observed between sodium excretion and CV events. However, further relevant studies are needed to validate our conclusions further.

#### AUTHOR CONTRIBUTIONS

ZHAO Dan analyzed the data and wrote the manuscript. LI Hua Min and LI Chao Xiu designed the tables and figures. ZHOU Bo proposed the idea for the study and supervised the whole study.

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