Letter





A Machine Learning-Based Prognostic Stratification Model for Liver Cancer: Results from Survival Studies Using SEER Database

Yuxuan Xiao^{1,2,3,&}, Zhuoying Li^{1,2,&}, Zhuojun Ye^{4,&}, Yu-Xin Zhou^{1,2,4}, Yixin Zou^{1,2,3}, Danni Yang^{1,2,4}, Yuting Tan^{1,2}, Qun Xu^{5,#}, and Yongbing Xiang^{1,2,3,4,5,#}

Primary liver cancer (PLC) is a major global health challenge, ranking as the sixth most common and third most fatal malignancy worldwide, according to GLOBOCAN 2022 estimates. [1] This high mortality rate underscores the aggressive nature of the disease and the significant burden it places on global healthcare systems. Although primary prevention remains the cornerstone of liver cancer control, improving outcomes for patients already diagnosed is equally critical for mitigating the impact of the disease. Currently, the international five-year survival rate for PLC is < 20%, [2] a stark figure that highlights the urgent and pressing need for more precise and reliable prognostic tools. The ability to accurately stratify patients based on their individual risks can directly guide personalized treatment decisions, allowing clinicians to tailor therapies to maximize efficacy while minimizing unnecessary toxicity. Such tools are essential for optimizing therapeutic strategies and preventing overtreatment in low-risk patients undertreatment in high-risk individuals, thereby helping control healthcare costs. Furthermore, robust prognostic models can accelerate innovative therapies development by rapidly identifying highrisk populations that are most likely to benefit from inclusion in randomized controlled trials.

Unfortunately, many of the prognostic systems currently in clinical use, such as the Barcelona Clinic Liver Cancer (BCLC), Okuda, and tumor-nodemetastasis (TNM) staging systems, are based on methodological frameworks that are now considered outdated. These systems often lack contemporary statistical validation and have significant limitations, such as poor discriminatory capacity for patients

with early-stage disease or being validated primarily specific, often advanced-stage, populations. Although new prognostic models have emerged, they face challenges in clinical translation, including the complexity of acquiring the necessary predictive variables, poor applicability across heterogeneous patient populations, and insufficient discriminative power. In recent years, the exponential growth in large-scale medical and health data has created opportunities to overcome these Traditional statistical methods increasingly struggling to meet the demands of complex data analysis. However, machine learning, with its powerful data processing and analytical capabilities, has provided a new paradigm for predicting cancer-related mortality and enabling precise, individualized prevention and control.[3] Therefore, this study aimed to develop and internally validate prognostic models for 60-month PLCspecific mortality using both traditional statistical and machine-learning approaches.

this context, this study leveraged comprehensive population-based data from the U.S. Surveillance, Epidemiology, and End Results (SEER) database to analyze the prognostic factors influencing survival in patients with PLC. Our The primary objective was to apply sophisticated machine learning algorithms to construct a prognostic stratification model specifically tailored to the diverse patients with liver cancer represented in this extensive database. The ultimate aim was to provide a clinically relevant tool for prognostic stratification that could serve as an evidence-based foundation for prognostic assessment, treatment selection, and follow-up management. By facilitating

doi: 10.3967/bes2025.143

^{1.} State Key Laboratory of System Medicine for Cancer, Shanghai Cancer Institute, Renji Hospital, Shanghai Jiao Tong University School of Medicine, Shanghai 200127, China; 2. Department of Epidemiology, Shanghai Cancer Institute, Shanghai 200032, China; 3. School of Public Health, Fudan University, Shanghai 200032, China; 4. School of Public Health, Shanghai Jiao Tong University School of Medicine, Shanghai 200025, China; 5. Health Management Center, Renji Hospital, Shanghai Jiao Tong University School of Medicine, Shanghai 200127, China

better-informed clinical decision-making, it is hoped that this will contribute to improved survival outcomes in patients with liver cancer.

To achieve this, data were extracted from 57,526 patients with PLC diagnosed between 2000 and 2017, with a complete follow-up period of at least 60 months. This study first employed both fine-gray competing risk and Cox proportional hazards regression models to identify key prognostic factors. Subsequently, the entire patient cohort was partitioned into training and test sets, upon which multiple machine-learning models were developed and rigorously compared. The best-performing model was then selected to build an individualized, user-friendly prediction tool for 60-month PLCspecific mortality risk. The variables analyzed in the study were: age (numeric); survival months (numeric); race (white and hispanic, black, Asian or Pacific Islander, American Indian/Alaska Native); year of diagnosis (2000-2005, 2006-2011, 2012-2017); median household income (inflation-adjusted to 2022): ≤90,000/year (household), >90,000/year (household); residential status (counties in metropolitan areas, counties not in metropolitan areas); marital status (married categories, single categories, unknown); histologic type [hepatocellular carcinoma (HCC), intrahepatic cholangiocarcinoma (ICC), other]; stage (localized, regional, distant, unknown/unstaged); surgery (no/unknown, yes); radiotherapy (no/unknown, yes); chemotherapy (no/unknown, yes); time from diagnosis to treatment (≤ 2 months, > 2 months, no/unknown); outcome [alive, dead (attributable to PLC), dead (attributable to other causes)]. The Appendix provides the detailed research methodology, including variable definitions, a full description of the statistical analysis, and the TRIPOD Checklist. The specific coding for each group is shown in eTable 1, where the group with the smallest code is considered the reference group.

RESULTS

This study cohort of 57,526 participants had a median age of 64 years and a median survival time of only 10 months, with a significant male predominance (73.66%). By the end of the follow-up period, 65.9% of these patients had died specifically from PLC within 60 months of diagnosis, confirming the high lethality of the disease. The baseline characteristics of the study population are presented in Table 1.

In the comprehensive analysis of prognostic factors, both the Fine-Gray competing risks model was utilized, accounting for events that preclude the event of interest (i.e., death from other causes), and the more traditional Cox proportional hazard regression model. The remarkable consistency between the results of the two models confirmed the robustness and reliability of the identified prognostic factors. The Fine-Gray model results (Table 2) revealed several demographic and clinical factors that significantly influenced PLC-specific survival.

In both the univariate and multivariable Cox regression analyses, the hazard ratios (HRs) and adjusted hazard ratios (aHRs) of all factors were statistically significant, except for the "American Indian/Alaska Native" category in the factor of race. Cox regression analysis yielded similar results to the Fine-Gray analysis, and it was observed that other histological types, compared to HCC, were also prognosis factors for PLC survival. Detailed results are shown in Table 3.

Consistent with previous research, older age, male sex, and black race were associated with a poorer prognosis. The elevated mortality risk among male patients may be linked to a higher prevalence of risk factors such as alcoholic liver disease and viral hepatitis, [4] including distinct molecular signatures. [5] The disparity faced by black patients, who have a

Table 1. The baseline	characteristics of SEER	patients with PLC	included in the study
-----------------------	-------------------------	-------------------	-----------------------

Characteristics	Overall	Alive	Dead (attributable to PLC)Dead (attributable to other causes)		
	(n =57,526)	(n =8,521)	(n =39,994)	(n =9,011)	
Survival (months), Median [Q1,	Q3] 10.00 [2.00, 41.00] 9	6.00 [64.00, 148.00]	6.00 [1.00, 19.00]	15.00 [3.00, 49.00]	
Age (years), Median [Q1, Q3]	64.00 [56.00, 73.00] !	59.00 [51.00, 66.00]	65.00 [57.00, 74.00]	65.00 [57.00, 75.00]	
Diagnosis, n (%)					
2000-2005	14,818 (25.76)	1,153 (13.53)	11,168 (27.92)	2,497 (27.71)	
2006-2011	20,335 (35.35)	2,543 (29.84)	14,415 (36.04)	3,377 (37.48)	
2012-2017	22,373 (38.89)	4,825 (56.62)	14,411 (36.03)	3,137 (34.81)	

Continued

Characteristics	Overall	Alive	Dead (attributable to PLC)	Dead (attributable to other causes)	
Characteristics	(n =57,526)	(n =8,521)	(n =39,994)	(n =9,011)	
Sex, n (%)					
Female	15,153 (26.34)	2,595 (30.45)	10,302 (25.76)	2,256 (25.04)	
Male	42,373 (73.66)	5,926 (69.55)	29,692 (74.24)	6,755 (74.96)	
Race, n (%)					
White and Hispanic	40,681 (70.72)	5,646 (66.26)	28,620 (71.56)	6,415 (71.19)	
Black	6,771 (11.77)	745 (8.74)	4,928 (12.32)	1,098 (12.19)	
Asian or Pacific Islander	9,524 (16.56)	2,064 (24.22)	6,046 (15.12)	1,414 (15.69)	
American Indian/Alaska Native	550 (0.96)	66 (0.77)	400 (1.00)	84 (0.93)	
Income (RMB/year, household), n (%)					
≤90,000	41865 (72.78)	5,720 (67.13)	29,599 (74.01)	6,546 (72.64)	
>90,000/year	15661 (27.22)	2,801 (32.87)	10,395 (25.99)	2,465 (27.36)	
Residential Status, n (%)					
Counties in metropolitan areas	51,768 (89.99)	7,956 (93.37)	35,632 (89.09)	8,180 (90.78)	
Counties not in metropolitan areas	5,758 (10.01)	565 (6.63)	4,362 (10.91)	831 (9.22)	
Marital status, n (%)					
Married categories	30,851 (53.63)	4,946 (58.04)	21,179 (52.96)	4,726 (52.45)	
Single categories	24,015 (41.75)	3,201 (37.57)	17,002 (42.51)	3,812 (42.30)	
Unknown	2,660 (4.62)	374 (4.39)	1,813 (4.53)	473 (5.25)	
Histologic Type, n (%)					
HCC	48,590 (84.47)	7,177 (84.23)	33,561 (83.92)	7,852 (87.14)	
ICC	1,723 (3.00)	60 (0.70)	1,444 (3.61)	219 (2.43)	
Other	7,213 (12.54)	1,284 (15.07)	4,989 (12.47)	940 (10.43)	
Stage, <i>n</i> (%)					
Localized	27,050 (47.02)	6,367 (74.72)	15,314 (38.29)	5,369 (59.58)	
Regional	14,487 (25.18)	1,408 (16.52)	11,220 (28.05)	1,859 (20.63)	
Distant	10,621 (18.46)	376 (4.41)	9,239 (23.10)	1,006 (11.16)	
Unknown/unstaged	5,368 (9.33)	370 (4.34)	4,221 (10.55)	777 (8.62)	
Surgery, n (%)					
No/unknown	38,694 (67.26)	1,950 (22.88)	31,423 (78.57)	5,321 (59.05)	
Yes	18,832 (32.74)	6,571 (77.12)	8,571 (21.43)	3,690 (40.95)	
Radiotherapy, n (%)					
No/Unknown	53,064 (92.24)	8,101 (95.07)	36,467 (91.18)	8,496 (94.28)	
Yes	4,462 (7.76)	420 (4.93)	3,527 (8.82)	515 (5.72)	
Chemotherapy, n (%)					
No/unknown	37,794 (65.70)	5,342 (62.69)	26,045 (65.12)	6,407 (71.10)	
Yes	19,732 (34.30)	3,179 (37.31)	13,949 (34.88)	2,604 (28.90)	
Time from diagnosis to treatment, n (%)					
≤ 2 months	21,255 (36.95)	4,747 (55.71)	13,236 (33.09)	3,272 (36.31)	
> 2 months	10,601 (18.43)	2,348 (27.56)	6,320 (15.80)	1,933 (21.45)	
No/Unknown	25,670 (44.62)	1,426 (16.74)	20,438 (51.10)	3,806 (42.24)	

higher mortality risk than White and Hispanic individuals, is a critical finding that likely reflects a complex interplay of socioeconomic differences, insufficient access to healthcare resources, and

potentially race-specific tumor biological behaviors. Conversely, Asian or Pacific Islander patients demonstrate significantly better prognoses, a finding that warrants further investigation into potential

Table 2. Univariate and multivariate Fine-Gray analysis of factors influencing survival in patients with PLC

Characteristics	Univariate analysis	- Р	Multivariable analysis sHR (95%CI)	- Р
Characteristics	sHR (95%CI)			
Age (per 1 year)	1.011 (1.011-1.011) ^a	< 0.001	1.008 (1.007-1.009) ^a	< 0.001
Sex				
Male vs. Female	1.050 (1.030-1.080) ^a	< 0.001	1.053 (1.027-1.079) ^a	< 0.001
Diagnosis year				
2006–2011 vs. 2000–2005	1.010 (0.990-1.030)	0.470	0.902 (0.878-0.927) ^b	< 0.001
2012–2017 vs. 2000–2005	0.843 (0.822-0.854) ^b	< 0.001	0.802 (0.781-0.825) ^b	< 0.001
Race				
Black vs. White and Hispanic	1.124 (1.092-1.166) ^a	< 0.001	1.044 (1.009-1.079) ^a	0.013
Asian or Pacific Islander vs. White and Hispanic	0.817 (0.795-0.839) ^b	< 0.001	0.916 (0.889-0.943) ^b	< 0.001
American Indian/Alaska Native vs. White and Hispanic	1.082 (0.980-1.184)	0.130	1.041 (0.942-1.150)	0.430
Income (RMB/year, household)				
>90,000 vs. ≤90,000	0.873 (0.852-0.894) ^b	< 0.001	0.968 (0.945-0.992) ^b	0.010
Residential Status				
Counties not in metropolitan areas vs. Counties in metropolitan areas	1.222 (1.182-1.252) ^a	< 0.001	1.126 (1.088-1.166) ^a	< 0.001
Marital status				
Single categories vs. Married categories	1.102 (1.071-1.123) ^a	< 0.001	1.025 (1.002-1.048) ^a	0.034
Unknown vs. Married categories	0.993 (0.953-1.043)	0.770	0.875 (0.830-0.922) ^b	< 0.001
Histologic Type				
ICC vs. HCC	1.742 (1.652-1.842) ^a	< 0.001	1.205 (1.133-1.283) ^a	< 0.001
Other vs. HCC	1.051 (1.021-1.082) ^a	0.001	0.996 (0.963-1.032)	0.840
Stage				
Regional vs. Localized	1.342 (1.323-1.371) ^a	< 0.001	1.602 (1.562-1.643) ^a	< 0.001
Distant vs. Localized	2.314 (2.254–2.373) ^a	< 0.001	2.123 (2.057-2.191) ^a	< 0.001
Unknown/unstaged vs. Localized	1.402 (1.361-1.452) ^a	< 0.001	1.294 (1.246-1.345) ^a	< 0.001
Surgery				
Yes vs. No/Unknown	0.303 (0.293-0.313) ^b	< 0.001	0.408 (0.396-0.420) ^b	< 0.001
Radiotherapy				
Yes vs. No/Unknown	1.172 (1.142-1.201) ^a	< 0.001	0.946 (0.914-0.978) ^b	0.001
Chemotherapy				
Yes vs. No/Unknown	0.896 (0.875-0.906) ^b	< 0.001	0.885 (0.863-0.907) ^b	< 0.001
Time from diagnosis to treatment				
>2 months vs. < = 2 months	0.593 (0.584-0.612) ^b	< 0.001	0.805 (0.784-0.826) ^b	< 0.001
No/Unknown vs. < = 2 months	2.131 (2.082-2.171) ^a	< 0.001	1.117 (1.083-1.151) ^a	< 0.001

Note. a, the interval estimate exceeds 1;b, the interval estimate is below 1. PLC, primary liver cancer; HCC, hepatocellular carcinoma; ICC, intrahepatic cholangiocarcinoma; sHR, sub-hazard ratio.

genetic, lifestyle, and healthcare-related protective factors^[6]. Socioeconomic status has emerged as a powerful determinant of outcomes. Household income exceeding \$90,000 per year was a significant

protective factor, highlighting the crucial role economic stability plays in comprehensive disease management. Patients with higher incomes are more likely to benefit from a timely diagnosis,

Table 3. Univariate and multivariate Cox regression analysis of factors influencing survival in patients with PLC

Characteristics	Univariate analysis	P	Multivariable analysis	P
	HR (95%CI)		aHR (95%CI)	
Age (per 1 year)	1.019 (1.018-1.019) ^a	< 0.001	1.014 (1.013-1.015) ^a	< 0.001
Sex				
Male vs. Female	1.075 (1.051-1.099) ^a	< 0.001	1.128 (1.102-1.155) ^a	< 0.001
Diagnosis year				
2006–2011 vs. 2000–2005	0.832 (0.812-0.853) ^b	< 0.001	0.892 (0.870-0.915) ^b	< 0.001
2012–2017 vs. 2000–2005	0.741 (0.723-0.760) ^b	< 0.001	0.793 (0.773-0.813) ^b	< 0.001
Race				
Black vs. White and Hispanic	1.135 (1.101-1.169) ^a	< 0.001	1.108 (1.075-1.143) ^a	< 0.001
Asian or Pacific Islander vs. White and Hispanic	0.784 (0.763-0.806) ^b	< 0.001	0.891 (0.866-0.917) ^b	< 0.001
American Indian/Alaska Native vs. White and Hispanic	1.038 (0.940-1.146)	0.459	1.016 (0.920-1.122)	0.756
Income (RMB/year, household)				
>90,000/year(Household) vs. <90,000/year(Household)	0.834 (0.815-0.853) ^b	< 0.001	0.941 (0.920-0.963) ^b	< 0.001
Residential Status				
Counties not in metropolitan areas vs. Counties in metropolitan areas	1.247 (1.209-1.287) ^a	< 0.001	1.137 (1.101-1.175) ^a	< 0.001
Marital status				
Single categories vs. Married categories	1.153 (1.130-1.177) ^a	< 0.001	1.095 (1.072-1.118) ^a	< 0.001
Unknown vs. Married categories	1.103 (1.052-1.157) ^a	< 0.001	0.883 (0.841-0.927) ^b	< 0.001
Histologic Type				
ICC vs. HCC	2.058 (1.952-2.170) ^a	< 0.001	1.311 (1.243-1.383) ^a	< 0.001
Other vs. HCC	1.077 (1.045-1.109) ^a	< 0.001	1.045 (1.013-1.077) ^a	0.005
Stage				
Regional vs. Localized	2.133 (2.081-2.186) ^a	< 0.001	1.807 (1.762-1.853) ^a	< 0.001
Distant vs. Localized	4.053 (3.947-4.163) ^a	< 0.001	2.784 (2.706-2.864) ^a	< 0.001
Unknown/unstaged vs. Localized	2.480 (2.397-2.567) ^a	< 0.001	1.263 (1.218-1.309) ^a	< 0.001
Surgery				
Yes vs. No/Unknown	0.215 (0.210-0.220) ^b	< 0.001	0.299 (0.289-0.308) ^b	< 0.001
Radiotherapy				
Yes vs. No/Unknown	1.144 (1.105-1.184) ^a	< 0.001	0.828 (0.797-0.860) ^b	< 0.001
Chemotherapy				
Yes vs. No/Unknown	0.808 (0.791-0.824) ^b	< 0.001	0.750 (0.730-0.770) ^b	< 0.001
Time from diagnosis to treatment				
>2 months vs. < = 2 months	0.837 (0.820-0.862) ^b	< 0.001	0.747 (0.724-0.770) ^b	< 0.001
No/Unknown vs. < = 2 months	2.622 (2.564-2.681) ^a	< 0.001	1.235 (1.197-1.274) ^a	< 0.001

Note. a, the interval estimate exceeds 1;b, the interval estimate is below 1. PLC, primary liver cancer; HCC, hepatocellular carcinoma; ICC, intrahepatic cholangiocarcinoma; HR, hazard ratio; aHR, adjusted hazard ratio.

standardized treatment protocols, and regular follow-up care, all of which contribute to improved survival. Furthermore, residing in a non-metropolitan area is a risk factor, likely reflecting disparities in the availability of specialized diagnostic and treatment facilities for liver cancer in rural regions.

Clinical- and treatment-related variables are strong predictors of survival. As expected, a more advanced tumor stage at diagnosis was strongly associated with increased mortality risk. Patients with ICC have a significantly higher risk of PLCspecific death than those with HCC, which is consistent with the aggressive biology of ICC, including its high invasiveness and poor response to systemic treatments. [8] Surgery, radiotherapy, and chemotherapy were all significantly associated with a reduced risk of death, with surgery demonstrating the most pronounced protective effect (sub-hazard ratio [sHR] = 0.408). These findings underscore the central and lifesaving role of curative-intent therapies in the management of patients with liver cancer. One of the more counterintuitive findings was that patients who began treatment > two months after diagnosis had a lower mortality risk. This is likely not because delaying treatment is beneficial, but rather an instance of "reverse causality." Patients who are in better overall physical condition may be more suitable candidates for a deliberative, multidisciplinary treatment planning process, which can take time. In contrast, patients with more advanced tumors or a poorer health often require immediate emergency treatment, thus creating a statistical artifact in which a longer diagnosis-to-treatment interval appears protective. [9] Cox regression analysis (Table 3) largely mirrored these results, further strengthening the validity of these factors as independent prognostic indicators for the overall survival of patients with liver cancer.

Leveraging these validated prognostic factors (with the exclusion of "year of diagnosis" to ensure the model's contemporary applicability), this study constructed and evaluated several machine learning models to predict the 60-month risk of PLC-specific death. The specific data for the training and test sets are listed in Table 2. The hyperparameter tuning results for each machine learning classifier are presented in Table 3. Among the various classifiers tested, the LightGBM and XGBoost models demonstrated superior predictive performances, achieving an impressive area under the receiver operating characteristic curve (AUC) of 0.770 (95%

CI: 0.762-0.777) on the independent test set (Figure 1).

The AUC is a measure of the model's ability to correctly distinguish between patients who will experience an event (death from PLC) and those who will not, with a value of 0.770 indicating a good level of discrimination. Given its comparable discrimination, similar calibration (shown eFigure 2), faster training speed, and lower memory consumption, LightGBM was selected to develop the final prediction tool. Furthermore, the Shapley Additive extension (SHAP) method was applied to assess the impact of each variable on the performance of the LightGBM model, and the detailed results are shown in eFigure 3.

The optimized LightGBM model was applied to the entire patient cohort for risk stratification. By dividing patients into three distinct groups based on their predicted mortality probability (low, medium, and high risk), the model proved highly effective in separating patient outcomes (Figure 2).

These findings have significant clinical implications. Patients in the high-risk group had a 10.52-fold higher risk of dying from PLC within 60 months than those in the low-risk group. The medium-risk group showed a 3.84-fold increase in the risk. This stratification provides valuable information for clinical practice. For instance, patients identified as low risk might be candidates

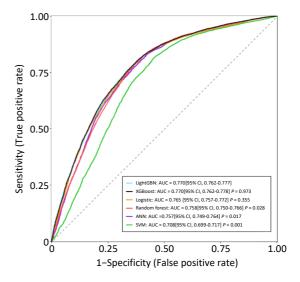


Figure 1. AUC of machine learning classifiers for predicting 60-month PLC-specific mortality on test set (ANN, artificial neural network; SVM, support vector machine; LightGBM, light gradient boosting machine; XGBoost, extreme gradient boosting machine).

for more conservative treatment strategies or less intensive surveillance, thereby avoiding the risks and costs of overtreatment. Conversely, patients stratified into the high-risk group should be prioritized for more aggressive therapeutic regimens, closer follow-up, and consideration for enrollment in clinical trials exploring novel treatment options. To illustrate, consider a 60-year-old, married, non-Hispanic White female residing in a metropolitan area with an annual household income of ≤\$90,000. She was diagnosed with localized HCC and treated within two months, without surgery, radiotherapy, or chemotherapy. This model calculated her 60-month PLC-specific mortality probability to be 0.5933, placing her in the mediumrisk category, which will help guide discussions on the intensity of her future care.

Despite the successful development of the highperformance prediction model, its limitations must be acknowledged. First, the SEER database, while large and comprehensive, lacks granular clinical data such as laboratory test results, [10] performance status, and detailed lifestyle factors. This lack of information may have led to the exclusion of important prognostic variables and introduced a residual confounding bias, potentially affecting the model's predictive accuracy at the individual level. Second, the model was developed and validated using a single, albeit large, population-based registration data set. However, this has not been validated in an external cohort from a different geographical region or healthcare system. Therefore, its generalizability and predictive performance in other patients or populations need to be examined and confirmed. Third, the model was specifically designed to predict the 60-month mortality.

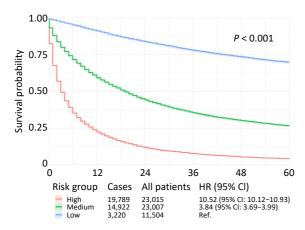


Figure 2. Survival curves for different risk groups in the entire patient cohort.

Although this is a standard endpoint, clinical scenarios often require prognostication over different time horizons (e.g., 1-year, 3-year, or 10-year survival). Future studies should focus on developing models for different follow-up durations to enhance clinical applicability. Finally, the intentional exclusion of the "year of diagnosis" variable, while aimed at improving the model's relevance to current practice, may have diminished its ability to capture the dynamic impact of recent therapeutic advancements, such as the widespread adoption of targeted therapies and immunotherapy, on PLC prognosis. Future iterations of the model should incorporate time-dependent variables to reflect the real-time impact of medical progress.

In conclusion, this study successfully leveraged the power of a large population-based dataset and advanced machine learning techniques to develop and validate a robust prognostic stratification tool for PLC. By integrating a set of readily accessible demographic and clinical variables, this tool provides an effective method for individualized risk assessment. This offers a novel and practical approach to support a more precise clinical prognosis and facilitate shared decision-making between physicians and patients, with the ultimate goal of improving outcomes for individuals affected by this challenging disease.

Funding National Key Project of Research and Development Program of China [2021YFC2500404]. Competing Interests None declared.

Ethics This study used publicly available deidentified data from the SEER database. According to the policies of the SEER program and relevant ethical guidelines, the use of such data does not require institutional review board (IRB) approval or informed consent.

Authors' Contributions The research was designed and funding obtained by Yongbing Xiang. The study was conducted by Yuxuan Xiao, Zhuoying Li, Zhuojun Ye, Yuxin Zhou, Yixin Zou, Danni Yang, Yuting Tan, Qun Xu, and Yongbing Xiang. Data analysis and interpretation were performed by Yu-Xuan Xiao, Zhuoying Li, Zhuojun Ye, and Yongbing Xiang. The first draft of the manuscript was written by Yuxuan Xiao, Qun Xu, and Yongbing Xiang. All authors reviewed and approved the final version of the manuscript. Yong-Bing Xiang had primary responsibility for the final content.

Acknowledgements We acknowledge the Surveillance, Epidemiology, and End Results (SEER) Program of the National Cancer Institute for

providing the data used in this study. We also extend our gratitude to patients whose data were included in the SEER database, which made this study possible. This work was supported by the National Key Project of the Research and Development Program of China [grant number 2021YFC2500404]. Data Sharing Publicly available datasets were analyzed in this study. Data were accessed and extracted using SEER*Stat software (version 8.4.4) from the Surveillance, Epidemiology, and End Results (SEER) database. The Supplementary Materials will be available in www.besjournal.com.

[&]These authors contributed equally to this work.

**Correspondence should be addressed to Qun Xu, E-mail: xuqun@renji.com; Yongbing Xiang, E-mail: ybxiang@shsci.org

Biographical notes of the first authors: Yuxuan Xiao, E-mail: yuxuanxiao7@foxmail.com; Zhuoying Li, E-mail: lzy2727@foxmail.com; Zhuojun Ye, E-mail: yezj1221@foxmail.com

Received: August 1, 2025; Accepted: October 16, 2025

REFERENCES

 Bray F, Laversanne M, Sung H, et al. Global cancer statistics 2022: GLOBOCAN estimates of incidence and mortality

- worldwide for 36 cancers in 185 countries. CA Cancer J Clin, 2024: 74, 229–63.
- Jiang YF, Li ZY, Ji XW, et al. Global pattern and trend of liver cancer survival: a systematic review of population-based studies. Hepatoma Res, 2020; 6, 52.
- Chakraborty A, Pant MD. Machine learning models for pancreatic cancer survival prediction: a multi-model analysis across stages and treatments using the surveillance, epidemiology, and end results (seer) database. J Clin Med, 2025: 14, 4686.
- Rich NE, Murphy CC, Yopp AC, et al. Sex disparities in presentation and prognosis of 1110 patients with hepatocellular carcinoma. Aliment Pharmacol Ther, 2020; 52, 701–09.
- Kim SY, Song HK, Lee SK, et al. Sex-biased molecular signature for overall survival of liver cancer patients. Biomol Ther (Seoul), 2020; 28, 491–502.
- Iheanacho F, Tramontano A, Manz C. Racial disparities in hepatocellular carcinoma (HCC) treatment and survival: a SEER-medicare analysis. JCO Oncol Pract, 2024; 20, 144.
- Rich NE, Jones PD, Zhu H, et al. Impact of racial, ethnic, and socioeconomic disparities on presentation and survival of HCC: a multicenter study. Hepatol Commun, 2024; 8, e0477.
- Penzkofer L, Gröger LK, Hoppe-Lotichius M, et al. Mixed hepatocellular cholangiocarcinoma: a comparison of survival between mixed tumors, intrahepatic cholangiocarcinoma and hepatocellular carcinoma from a single center. Cancers (Basel), 2023; 15, 639.
- Wei TT, Huang HB, Zhang AJ, et al. Impact of the diagnosis-totreatment interval on the survival of patients with papillary thyroid cancer. J Invest Surg, 2025; 38, 2456463.
- Gao Y, Liu J, Zhao DX, et al. A novel prognostic model for identifying the risk of hepatocellular carcinoma based on angiogenesis factors. Front Genet, 2022; 13, 857215.